

# Analytic Modeling

## Birth-Death Model

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## A Review -Random Variables

- A variable representing the outcome of a random activity, e.g., rolling a die
- Defines a mapping from outcome to value
- Has a range of values over which it can vary and a probability distribution with which it takes on these values
- Discrete random variable -can take on a finite or countable set of values, e.g., number of customers in system
- Continuous random variable -can take on values over a continuous interval, e.g., waiting time

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## Discrete Random Variables

- Discrete probability function,  $P(n)$

$$P(n) = \Pr[\tilde{x} = n]$$

where  $0 \leq P(n) \leq 1$  and  $\sum_n P(n) = 1$

- Cumulative distribution function,  $F(x)$

$$F(x) = \Pr[\tilde{x} \leq x] = \sum_{n \leq x} P(n)$$

where  $F(0) = 0$

$$F(\infty) = 1$$

$$F(b) \geq F(a) \text{ if } b \geq a$$

$$F(b) - F(a) = \Pr[a < \tilde{x} < b] = \sum_{a < n < b} P(n)$$

## Discrete Random Variables

- Mean and variance

$$\text{mean: } \bar{x} = \sum_n nP(n)$$

$$\text{variance: } \sigma^2 = \sum_n (n - \bar{x})^2 P(n)$$

## Continuous Random Variables

- Probability density function,  $f(x)$

$$f(x) \geq 0$$

$$\Pr[a \leq \tilde{x} \leq b] = \int_a^b f(x) dx$$

$$\int_0^{\infty} f(x) dx = 1$$

- Cumulative distribution function,  $F(x)$

$$F(x) = \Pr[\tilde{x} \leq x] = \int_0^x f(y) dy$$

$$f(x) = \frac{dF(x)}{dx}$$

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## Continuous Random Variables

- Mean and variance

$$\text{mean: } \bar{x} = \int_0^{\infty} xf(x) dx$$

$$\text{variance: } \sigma^2 = \int_0^{\infty} (x - \bar{x})^2 f(x) dx$$

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## Exponential Distribution (with Parameter $\lambda$ )

$$\circ f(x) = \lambda e^{-\lambda x} \quad x \geq 0$$

$$F(x) = 1 - e^{-\lambda x}$$

$$\bar{x} = \frac{1}{\lambda}, \quad \sigma^2 = \frac{1}{\lambda^2}$$

○ Has memoryless property

○ Coefficient of variation =  $\frac{\sigma}{\bar{x}} = 1$

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## Memoryless Property

$$\circ \Pr[x_0 < \tilde{x} \leq x + x_0 \mid \tilde{x} > x_0]$$

$$= \frac{\Pr[x_0 < \tilde{x} < x + x_0]}{\Pr[\tilde{x} > x_0]}$$

$$= \frac{F(x + x_0) - F(x_0)}{1 - F(x_0)}$$

$$= \frac{1 - e^{-\lambda(x+x_0)} - 1 + e^{-\lambda x_0}}{1 - 1 + e^{-\lambda x_0}}$$

$$= 1 - e^{-\lambda x}$$

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## Memoryless Property -Interpretation

- Suppose interarrival time has exponential distribution with parameter  $\lambda$
- Suppose an arrival occurs at time 0
- Given that there is no arrival between 0 and  $x_0$ , the time until the next arrival, measured from  $x_0$ , has exponential distribution with parameter  $\lambda$
- Since time 0 and  $x_0$  are arbitrarily chosen, the time until the next arrival, measured from any time instant, has exponential distribution with parameter  $\lambda$
- *Note:*  $1/\lambda$  is the mean interarrival time, and  $\lambda$  is the arrival rate

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## More on Exponential Distribution

- Suppose interarrival time has exponential distribution with parameter  $\lambda$

$$\Pr[\text{no arrival in } (0, h)]$$

$$= 1 - F(h)$$

$$= e^{-\lambda h}$$

$$= 1 - \lambda h + o(h)$$

$$\text{where } o(h) \text{ has the property } \lim_{h \rightarrow 0} \frac{o(h)}{h} = 0$$

(an order of magnitude smaller than  $h$ )

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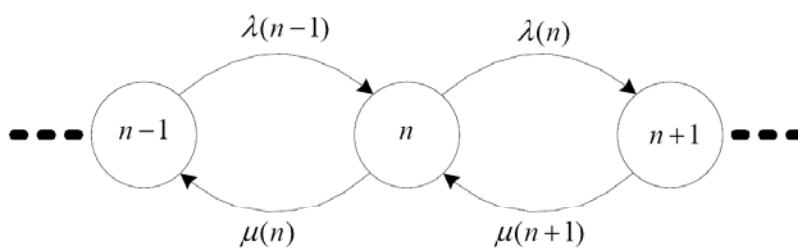
## More on Exponential Distribution

$$\begin{aligned}\Pr[1 \text{ arrival in } (0, h)] &= \int_0^h \lambda e^{-\lambda y} dy \cdot e^{-\lambda(h-y)} \\ &= \lambda h e^{-\lambda h} \\ &= \lambda h + o(h)\end{aligned}$$

$$\begin{aligned}\Pr[2 \text{ or more arrivals in } (0, h)] &= 1 - \Pr[\text{no arrival in } (0, h)] - \Pr[1 \text{ arrival in } (0, h)] \\ &= o(h)\end{aligned}$$

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## Birth-Death Model



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## Birth-Death Model

- Queuing system with a single service facility
- Interarrival time is exponentially distributed
- Service time is exponentially distributed
- Arrival rate and service rate may be state dependent, i.e., a function of the state of the system

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## State Dependent Arrival Rate

- While the state of the system is  $j$ , interarrival time has exponential distribution with parameter  $\lambda(j)$ , e.g.,  $\lambda(j) = (N - j)\lambda$ 
  - $\lambda(j)$  is also the arrival rate
  - if state is changed, say to  $k$ , the arrival rate (or the parameter of exponential distribution) is immediately changed to  $\lambda(k)$
- Useful for modeling the following features
  - finite population
  - finite queueing space

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## State Dependent Service Rate

- While the state of the system is  $j$ , service time has exponential distribution with parameter  $\mu(j)$ , e.g.,  $\mu(j) = j\mu$ 
  - $\mu(j)$  is also the service rate
  - if state is changed, say to  $k$ , the service rate (or the parameter of exponential distribution) is immediately changed to  $\mu(k)$
- Useful for modeling the following features
  - multiple servers
  - infinite servers

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## Definition of Birth-Death Process

- Stochastic process  $\tilde{n}(t)$  - a series of random variables indexed by time  $t$
- Let  $\tilde{n}(t)$  be the no. of customers in system at time  $t$ , and
- $q_{j,k}(h) = \Pr[\tilde{n}(t+h) = k \mid \tilde{n}(t) = j]$ , which is called transition probability
- Definition  $\tilde{n}(t)$  is a birth-death process if
  - i.  $q_{j,j+1}(h) = \lambda(j)h + o(h)$
  - ii.  $q_{j,j-1}(h) = \mu(j)h + o(h)$
  - iii.  $q_{j,j}(h) = 1 - \lambda(j)h - \mu(j)h + o(h)$
  - iv.  $q_{j,k}(h) = o(h)$  for  $|j-k| > 1$

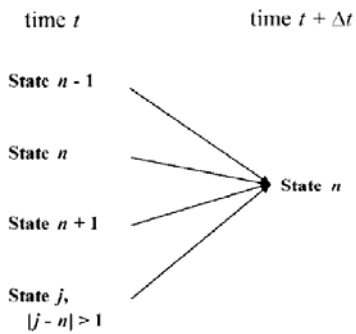
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## State Transition

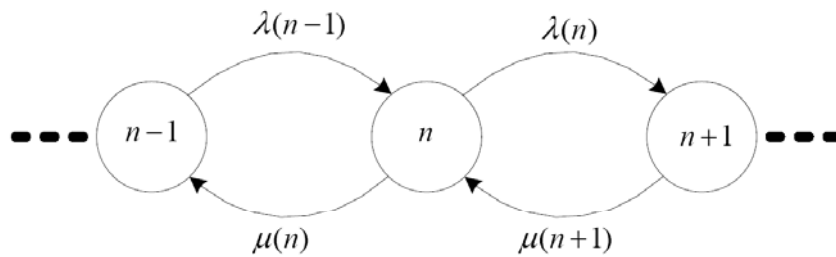
○ Let  $P(n, t) = \Pr[\tilde{n}(t) = n]$

Consider  $n > 1$



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## State Transitions



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# Analysis

$$\begin{aligned}P(n, t + \Delta t) &= P(n-1, t)[\lambda(n-1)\Delta t + o(\Delta t)] \\ &\quad + P(n, t)[1 - \lambda(n)\Delta t - \mu(n)\Delta t + o(\Delta t)] \\ &\quad + P(n+1, t)[\mu(n+1)\Delta t + o(\Delta t)] + o(\Delta t)\end{aligned}$$

Rearranging and dividing by  $\Delta t$ ,

$$\begin{aligned}\frac{P(n, t + \Delta t) - P(n, t)}{\Delta t} &= \lambda(n-1)P(n-1, t) - [\lambda(n) + \mu(n)]P(n, t) \\ &\quad + \mu(n+1)P(n+1, t) + \frac{o(\Delta t)}{\Delta t}\end{aligned}$$

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As  $\Delta t \rightarrow 0$ , we get, for  $n > 0$

$$\begin{aligned}\frac{\partial P(n, t)}{\partial t} &= \lambda(n-1)P(n-1, t) - [\lambda(n) + \mu(n)]P(n, t) \\ &\quad + \mu(n+1)P(n+1, t)\end{aligned}$$

For  $n = 1$

$$\frac{\partial P(0, t)}{\partial t} = -\lambda(0)P(0, t) + \mu(1)P(1, t)$$

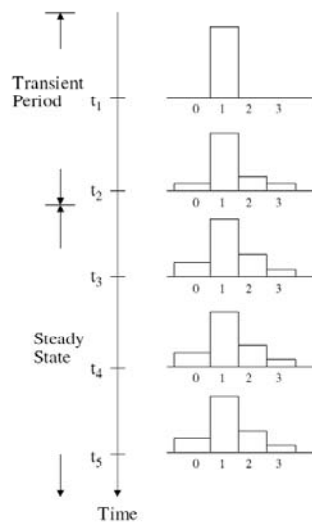
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## Remarks

- Analysis leads to a set of differential-difference equations which are difficult to solve in general
- Normal practice is to obtain steady state results for  $P(n, t)$  (i.e.,  $P(n, t)$  as  $t \rightarrow \infty$ )
  - steady state results are much easier to obtain
  - useful if one is interested in the long-term behavior of the system
  - usually considered as adequate when high-level performance estimates are required

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## Steady State Behavior



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## Steady State Probability

- Define steady state probability distribution

$$P(n) = \lim_{t \rightarrow \infty} P(n, t) \quad \Rightarrow \quad \frac{\partial P(n, t)}{\partial t} = 0$$

- $P(n)$  is the long-run probability that there are  $n$  customers in the system

- We have a set of *balance equations* (at steady state):

For  $n > 0$ ,

$$[\lambda(n) + \mu(n)]P(n) = \lambda(n-1)P(n-1) + \mu(n+1)P(n+1)$$

For  $n = 0$ ,

$$\lambda(0)P(0) = \mu(1)P(1)$$

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- Solving the above equations

$$P(n+1) = \frac{\lambda(n)}{\mu(n+1)} P(n), \quad n \geq 0$$

$$\Rightarrow P(n) = P(0) \prod_{i=0}^{n-1} \frac{\lambda(i)}{\mu(i+1)}, \quad n \geq 0$$

Since  $\sum_n P(n) = 1$ , then

$$P(0) \sum_{n=0}^{\infty} \prod_{i=0}^{n-1} \frac{\lambda(i)}{\mu(i+1)} = 1$$

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## Normalization Constant, $G$

○ Let  $G = \sum_{n=0}^{\infty} \prod_{i=0}^{n-1} \frac{\lambda(i)}{\mu(i+1)}$

○ If the series  $G$  converges,  $P(0) = 1/G$ , steady state solution exists and

$$P(n) = P(0) \prod_{i=0}^{n-1} \frac{\lambda(i)}{\mu(i+1)}, \quad n \geq 0$$

○ Otherwise,  $P(n) = 0$ , for  $n = 0, 1, \dots$